

# Blocking in multi-stage experiments

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# Abstract

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Is it better to align the Stage 2 blocks with the Stage 1 blocks as  
far as possible  
or to make them as orthogonal to each other as possible?

In either case, how should treatments be assigned?

## Example 1: the problem

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How should we design the experiment?

## Example 1: design 1a

**Design 1a** Align batches with lots,  
and allocate the 6 combinations of levels of  $F$  and  $G$   
in a randomized complete-block design.

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Mean	1	Mean	1	$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2 + q_0$
Blocks	5			$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2$
Units[B] 30		$F$	1	$\sigma^2 + q(F)$
		$G$	2	$\sigma^2 + q(G)$
		$F\#G$	2	$\sigma^2 + q(FG)$
		residual	25	$\sigma^2$

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We have lost 5 residual degrees of freedom, and gained nothing.

# An easy lesson

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Brien, Harch, Correll and Bailey (2011) call this  
“confounding big with big”.

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		residual	4	$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2$
Units[B]	30	$G$	2	$\sigma^2 + q(G)$
		$F\#G$	2	$\sigma^2 + q(FG)$
		residual	26	$\sigma^2$

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The variance for the contrast between levels of  $F$  involves  $\sigma_L^2$  as well as  $\sigma_B^s$ , so it is larger than it needs to be.

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There are 4 residual degrees of freedom for testing the main effect of  $F$ , and this cannot be increased.

## Example 2: design 2b

**Design 2b** Cross batches with lots to form a square array;  
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and then randomize rows and columns.

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source	df	source	df	EMS
Mean	1	Mean	1	$6\sigma_B^2 + 6\sigma_L^2 + \sigma^2 + q_0$
Batches	5	$F$	1	$6\sigma_B^2 + \sigma^2 + q(F)$
		residual	4	$6\sigma_B^2 + \sigma^2$
Lots	5			$6\sigma_L^2 + \sigma^2$
B#L	25	$G$	2	$\sigma^2 + q(G)$
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		residual	21	$\sigma^2$

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Main effect of  $F$  has smaller variance than before, and same df.

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Main effect of  $F$  has smaller variance than before, and same df. Other treatment effects have same (small) variance, and df reduced from 26 to 21.

# Some principles

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*If a treatment factor has to be applied to large units such as blocks in one stage,  
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*If a treatment factor has to be applied to large units such as blocks in one stage,  
then it will have relatively few residual degrees of freedom.  
In order not to reduce these further, try to confound  
the whole of this block term with the same term in other stages.*

## Example 2: design 2c

**Design 2c** Make 3 squares by crossing pairs of batches (shown as rows) with pairs of lots (shown as columns).

<i>F1</i>	G1, G2, G3	G1, G2, G3
<i>F2</i>	G1, G2, G3	G1, G2, G3

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Randomize levels of *F* to rows within each square;  
randomize levels of *G* within each corner of each square.

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Randomize levels of *F* to rows within each square;  
randomize levels of *G* within each corner of each square.

**Remark** Squares is the **supremum** of Batches and Lots:  
 $\text{Squares} = \text{Batches} \vee \text{Lots}.$

## Example 2: skeleton anova for design 2c

units		treatments		
source	df	source	df	EMS
Mean	1	Mean	1	$6\sigma_B^2 + 6\sigma_L^2 + \sigma_1^2 + q_0$
Squares	2			$6\sigma_B^2 + 6\sigma_L^2 + \sigma_1^2$
Batches[S]	3	F	1	$6\sigma_B^2 + \sigma_1^2 + q(F)$
		residual	2	$6\sigma_B^2 + \sigma_1^2$
Lots[S]	3			$6\sigma_L^2 + \sigma_1^2$
B#L[S]	3			$\sigma_1^2$
Units[B,L,S]	24	G	2	$\sigma^2 + q(G)$
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		residual	20	$\sigma^2$

The randomization argument suggests that  $\sigma_1^2 \neq \sigma^2$ .

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Both residual df have decreased, and nothing has been gained.

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Any further blocking of either batches or lots reduces the already-small residual df for main effects.

## More general numbers

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with each intersection containing  $Nm/bc$  experimental units,  
where  $m$  divides  $b$ ,  $m$  divides  $c$  and  $bc$  divides  $Nm$ ,  
so  $m$  divides  $\gcd\{b, c\}$ , and  $\gcd\{b, c\}$  divides  $Nm/\text{lcm}\{b, c\}$ .

## Example 4: arrays

There are 36 experimental units.

In Stage 1, these must be processed in 12 batches of size 3.

In Stage 2, the units must be processed in 9 lots of size 4.

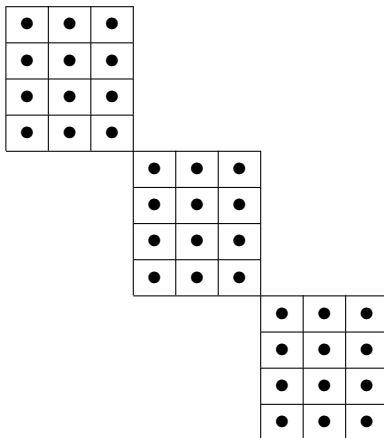
## Example 4: arrays

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In Stage 1, these must be processed in 12 batches of size 3.

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For orthogonality, the only possibility is three  $4 \times 3$  arrays.



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It is desirable to keep  $m$  small.

## Powers of the same small prime

Suppose that the number of batches and the number of lots are both powers of  $p$ , where  $p = 2$  or  $p = 3$ , that several  $p$ -level treatment factors  $F_1, F_2, \dots$  must be applied to whole batches in Stage 1, and several  $p$ -level treatment factors  $G_1, G_2, \dots$  must be applied to whole lots in Stage 2.

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If each stage is a single replicate of the relevant treatments, we may be able to take  $m = p$  and sacrifice information only on high-order interactions.

For example, if  $p = 2$  and  $F_1, F_2, F_3$  and  $F_4$  are applied to whole batches in Stage 1 while  $G_1, G_2$  and  $G_3$  are applied to whole lots in Stage 2, we can use two  $8 \times 4$  arrays, and confound  $F_1 F_2 F_3 F_4$  and  $G_1 G_2 G_3$  with each other and with arrays.

## Powers of the same small prime

Suppose that the number of batches and the number of lots are both powers of  $p$ , where  $p = 2$  or  $p = 3$ , that several  $p$ -level treatment factors  $F_1, F_2, \dots$  must be applied to whole batches in Stage 1, and several  $p$ -level treatment factors  $G_1, G_2, \dots$  must be applied to whole lots in Stage 2.

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This technique is called **post-fractionation** by Bisgaard (1997) and Vivacqua and Bisgaard (2009).

# Three stages

Suppose that there are three stages, and that treatment factors  $F$ ,  $G$  and  $H$  are applied in Stages 1, 2, 3 respectively.

Suppose that each treatment factor must be applied to whole blocks in its stage.

We already know that we should try to make the blocks from each stage as orthogonal as possible to blocks from every other stage.

An example was investigated by Mee and Bates (1998).

## Example 5: the problem

There are 16 experimental units.

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In Stage 1, these must be processed in 4 batches of size 4.

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There are 16 experimental units.

In Stage 1, these must be processed in 4 batches of size 4.  
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Treatment factor  $G$  has 2 levels, which are applied in Stage 2,  
and these must be applied to whole lots.

In Stage 3, the units must be processed in 4 pods of size 4.  
Treatment factor  $H$  has 2 levels, which are applied in Stage 3,  
and these must be applied to whole pods.

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In Stage 1, these must be processed in 4 batches of size 4.  
Treatment factor  $F$  has 2 levels, which are applied in Stage 1,  
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Treatment factor  $G$  has 2 levels, which are applied in Stage 2,  
and these must be applied to whole lots.

In Stage 3, the units must be processed in 4 pods of size 4.  
Treatment factor  $H$  has 2 levels, which are applied in Stage 3,  
and these must be applied to whole pods.

How should we design the experiment?

## Example 5: design 5a

**Design 5a** Form the experimental units into 2 arrays of size  $2 \times 2 \times 2$ . The first coordinate indicates the batch, the second coordinate indicates the lot, and the third coordinate indicates the pod. Within each array, randomize levels of  $F$  to batches, levels of  $G$  to lots, and levels of  $H$  to pods.

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The single df between Arrays is wasted.

All 7 treatment df have different variances, each with just one residual df.

## Example 5: design 5b

**Design 5b** Let  $\Lambda$  be a  $4 \times 4$  Latin square.

Stage 1: identify the batches with the rows of  $\Lambda$ .

Stage 2: identify the lots with the columns of  $\Lambda$ .

Stage 3: identify the pods with the letters  $\Lambda$ .

## Example 5: design 5b

**Design 5b** Let  $\Lambda$  be a  $4 \times 4$  Latin square.

Stage 1: identify the batches with the rows of  $\Lambda$ .

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Can we apply levels of  $F, G, H$  to rows, columns, letters respectively in such a way that all treatment interactions are orthogonal to rows, columns and letters?

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Can we apply levels of  $F, G, H$  to rows, columns, letters respectively in such a way that all treatment interactions are orthogonal to rows, columns and letters?

Mee and Bates found a cunning way of doing this, using the non-cyclic Latin square of order 4.

## Example 5: skeleton anova for design 5b

units		treatments		
source	df	source	df	EMS
Mean	1	Mean	1	$4\sigma_R^2 + 4\sigma_C^2 + 4\sigma_L^2 + \sigma^2 + q_0$
Rows	3	$F$	1	$4\sigma_R^2 + \sigma^2 + q(F)$
		residual	2	$4\sigma_R^2 + \sigma^2$
Columns	3	$G$	1	$4\sigma_C^2 + \sigma^2 + q(G)$
		residual	2	$4\sigma_C^2 + \sigma^2$
Letters	3	$H$	1	$4\sigma_L^2 + \sigma^2 + q(H)$
		residual	2	$4\sigma_L^2 + \sigma^2$
Units[R,C,L]	6	$F\#G$	1	$\sigma^2 + q(FG)$
		$F\#H$	1	$\sigma^2 + q(FH)$
		$G\#H$	1	$\sigma^2 + q(GH)$
		$F\#G\#H$	1	$\sigma^2 + q(FGH)$
		residual	2	$\sigma^2$

## Example 5: skeleton anova for design 5b

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source	df	source	df	EMS
Mean	1	Mean	1	$4\sigma_R^2 + 4\sigma_C^2 + 4\sigma_L^2 + \sigma^2 + q_0$
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		residual	2	$4\sigma_R^2 + \sigma^2$
Columns	3	$G$	1	$4\sigma_C^2 + \sigma^2 + q(G)$
		residual	2	$4\sigma_C^2 + \sigma^2$
Letters	3	$H$	1	$4\sigma_L^2 + \sigma^2 + q(H)$
		residual	2	$4\sigma_L^2 + \sigma^2$
Units[R,C,L]	6	$F\#G$	1	$\sigma^2 + q(FG)$
		$F\#H$	1	$\sigma^2 + q(FH)$
		$G\#H$	1	$\sigma^2 + q(GH)$
		$F\#G\#H$	1	$\sigma^2 + q(FGH)$
		residual	2	$\sigma^2$

For this Latin square, the decomposition into strata can be justified by randomization (Bailey, 1982).

## Example 5: design 5b by design key

The design key introduced by Patterson (1965) gives a clean construction of design 5b. All factors and pseudofactors have two levels, and arithmetic is modulo 2.

Rows:  $R_1, R_2, R_1 + R_2$ .

Columns:  $C_1, C_2, C_1 + C_2$ .

Letters:  $L_1 = R_1 + C_1, L_2 = R_2 + C_2, L_1 + L_2 = R_1 + R_2 + C_1 + C_2$ .

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$F = R_1$ .

$G = C_2$ .

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$F = R_1$ .

$G = C_2$ .

$H = R_1 + R_2 + C_1 + C_2$ .

$F + G = R_1 + C_2$ .

$F + H = R_2 + C_1 + C_2$ .

$G + H = R_1 + R_2 + C_1$ .

$F + G + H = R_2 + C_1$ .

# Non-orthogonality

So far, we have assumed that treatments applied in Stage  $i$   
**either** must be applied to whole blocks in Stage  $i$   
**or** can be orthogonal to blocks in Stage  $i$ .

Suppose that neither of these conditions holds?

## Example 6: the problem

There are 36 experimental units.

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Stage 1 has 12 blocks of size 3, in 4 superblocks of 3 blocks.

Treatment factor  $F$  has 3 levels, which are applied in Stage 1, and these must be applied to whole blocks.

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Stage 2 is a rectangle with 4 rows and 9 columns.

Treatment factor  $G$  has 3 levels, which are applied in Stage 2, in such a way that each combination of  $F$  and  $G$  occurs 4 times.

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Forced:  $F$  applied to whole blocks.

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Good idea: Apply each level of  $F$  to one block per superblock.

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Good idea: Align superblocks and rows.

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Good idea: Form the 12 blocks and 9 columns into three arrays of size  $4 \times 3$ .

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Good idea: Make the design in columns as efficient as possible.

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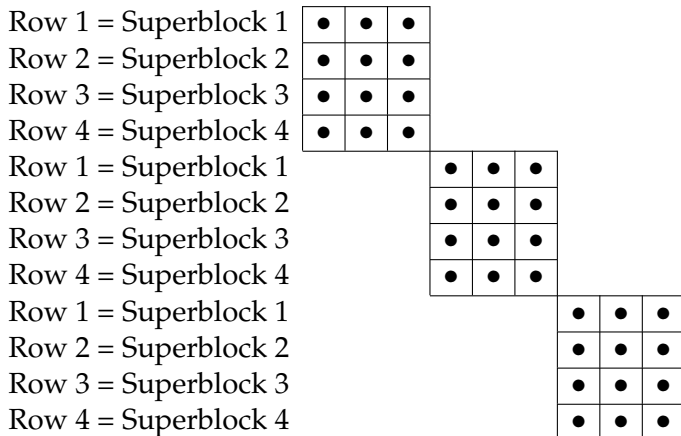
Good idea: Align superblocks and rows.

Good idea: Form the 12 blocks and 9 columns into three arrays of size  $4 \times 3$ .

Good idea: Make the design in columns as efficient as possible.

Can we achieve all of this?

## Example 6: align superblocks with rows, and make arrays



## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	F1G1	F1G2	F1G3	F2G1	F2G2	F2G3	F3G1	F3G2	F3G3

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	F1G1	F1G2	F1G3	F2G1	F2G2	F2G3	F3G1	F3G2	F3G3
Row 2	F2G3	F2G1	F2G2	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2
Row 3	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2	F2G3	F2G1	F2G2
Row 4	F1G2	F1G3	F1G1	F2G2	F2G3	F2G1	F3G2	F3G3	F3G1

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	F1G1	F1G2	F1G3	F2G1	F2G2	F2G3	F3G1	F3G2	F3G3
Row 2	F2G3	F2G1	F2G2	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2
Row 3	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2	F2G3	F2G1	F2G2
Row 4	F1G2	F1G3	F1G1	F2G2	F2G3	F2G1	F3G2	F3G3	F3G1

F1	F1	F1	F2	F2	F2	F3	F3	F3
G3	G1	G2	G3	G1	G2	G3	G1	G2

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	F1G1	F1G2	F1G3	F2G1	F2G2	F2G3	F3G1	F3G2	F3G3
Row 2	F2G3	F2G1	F2G2	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2
Row 3	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2	F2G3	F2G1	F2G2
Row 4	F1G2	F1G3	F1G1	F2G2	F2G3	F2G1	F3G2	F3G3	F3G1

F1	F1	F1	F2	F2	F2	F3	F3	F3
G3	G1	G2	G3	G1	G2	G3	G1	G2

## Example 6: design

	Array 1			Array 2			Array 3		
Row 1	F1G1	F1G2	F1G3	F2G1	F2G2	F2G3	F3G1	F3G2	F3G3
Row 2	F2G3	F2G1	F2G2	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2
Row 3	F3G3	F3G1	F3G2	F1G3	F1G1	F1G2	F2G3	F2G1	F2G2
Row 4	F1G2	F1G3	F1G1	F2G2	F2G3	F2G1	F3G2	F3G3	F3G1

F1	F1	F1	F2	F2	F2	F3	F3	F3
G3	G1	G2	G3	G1	G2	G3	G1	G2

The design in columns is factorially balanced,  
with canonical efficiency factors  $15/16$  for both main effects  
and  $3/4$  for the interaction.

## Example 6: skeleton anova

units		treatments			
source	df	cef	source	df	EMS
Mean	1	1	Mean	1	$9\sigma_R^2 + 9\sigma_S^2 + 4\sigma_C^2 + 3\sigma_B^2 + \sigma^2 + q_0$
Rows	3				$9\sigma_R^2 + 9\sigma_S^2 + 3\sigma_B^2 + \sigma^2$
Arrays	2	$\frac{1}{16}$	F	2	$4\sigma_C^2 + 3\sigma_B^2 + \sigma^2 + \frac{1}{16}q(F)$
Blocks[R,A]	6	$\frac{15}{16}$	F	2	$3\sigma_B^2 + \sigma^2 + \frac{15}{16}q(F)$
			residual	4	$3\sigma_B^2 + \sigma^2$
Columns[A]	6	$\frac{1}{16}$	G	2	$4\sigma_C^2 + \sigma^2 + \frac{1}{16}q(G)$
		$\frac{1}{4}$	F#G	4	$4\sigma_C^2 + \sigma^2 + \frac{1}{4}q(FG)$
Units[B,C,R,A]	18	$\frac{15}{16}$	G	2	$\sigma^2 + \frac{15}{16}q(G)$
		$\frac{3}{4}$	F#G	4	$\sigma^2 + \frac{3}{4}q(FG)$
			residual	12	$\sigma^2$

## A strategy for non-orthogonality

Suppose that, in Stage 1, treatment factor  $F$  has canonical efficiency factors

$$\begin{array}{ll} p & \text{in Batches} \\ q & \text{in Units[Batches]}, \end{array}$$

where  $p + q = 1$ .

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$$\begin{aligned} p & \text{ in Batches} \\ q & \text{ in Units[Batches]}, \end{aligned}$$

where  $p + q = 1$ .

Possible strategies for Stage 2:

- $(F \perp L)$  make  $F$  orthogonal to Lots;
- (BigwithBig) confound (the  $F$ -part of) Batches with Lots and the  $F$ -part of  $U[B]$  with  $U[L]$ ;
- (Compensate) confound the  $F$ -part of Batches with  $U[L]$  and the  $F$ -part of  $U[B]$  with Lots.

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If  $p = 1$  then  $F$  is confounded with Batches, so  $(F \perp L)$  is the same as (Compensate) and this is best.

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If  $q = 1$  then  $F$  is orthogonal to Batches, so  $(F \perp L)$  is the same as (BigwithBig) and this is best.

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If  $q = 1$  then  $F$  is orthogonal to Batches, so  $(F \perp L)$  is the same as (BigwithBig) and this is best.

For all values of  $p$ ,  $(F \perp L)$  is best if it is possible.

# A strategy for non-orthogonality, continued

## Theorem

*If Stage 1 has  $b$  batches of size  $s$  and Stage 2 has  $c$  lots of size  $k$ , then (BigwithBig) is better than (Compensate) in the sense of giving smaller variances for the estimators of contrasts between levels of  $F$  if and only if*

$$\frac{q}{p} > \frac{\sigma^2}{(s\sigma_B^2 + \sigma^2)} \frac{(k\sigma_L^2 + \sigma^2)}{(s\sigma_B^2 + k\sigma_L^2 + \sigma^2)}.$$

## A strategy for non-orthogonality, continued

### Theorem

*If Stage 1 has  $b$  batches of size  $s$  and Stage 2 has  $c$  lots of size  $k$ , then (BigwithBig) is better than (Compensate) in the sense of giving smaller variances for the estimators of contrasts between levels of  $F$  if and only if*

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If  $q > p$  then (BigwithBig) is better;  
otherwise, prior estimates of the relative magnitudes  
of  $\sigma^2$ ,  $\sigma_B^2$  and  $\sigma_L^2$  are required to make the decision.

## Another strategy for non-orthogonality

Stage 1 has  $b$  batches of size  $s$  and Stage 2 has  $c$  lots of size  $k$ ,  
so  $N = bs = ck$ .

If  $s$  divides  $k$  then we could nest batches within lots.

If  $sk$  divides  $N$  then we could cross batches with lots in  
 $N/sk$  arrays.

If  $k = ts$  and  $N = msk$  then we could do either.

## Another strategy for non-orthogonality

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so  $N = bs = ck$ .

If  $s$  divides  $k$  then we could nest batches within lots.

If  $sk$  divides  $N$  then we could cross batches with lots in  
 $N/sk$  arrays.

If  $k = ts$  and  $N = msk$  then we could do either.

Which is better?

## Another strategy for non-orthogonality, continued

Stage 1 has  $mts$  batches of size  $s$  and

Stage 2 has  $ms$  lots of size  $ts$ .

Let  $\Delta$  be a design for  $F$  in  $mts$  batches of size  $s$ .

Let  $\Gamma$  be a design for  $F$  in  $ms$  lots of size  $ts$ .

Let  $\text{Nest}(\Delta, \Gamma)$  be a design where each lot contains  $t$  batches, the design in batches is  $\Delta$  and the design in lots is  $\Gamma$ .

Let  $\text{Cross}(\Delta, \Gamma)$  be a design with  $m$  arrays of  $ts$  batches crossed with  $s$  lots, the design in batches is  $\Delta$  and the design in lots is  $\Gamma$ .

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### Theorem

*The variances for the estimators of contrasts between levels of  $F$  are no bigger for  $\text{Nest}(\Delta, \Gamma)$  than for  $\text{Cross}(\Delta, \Gamma)$ .*

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### Theorem

*The variances for the estimators of contrasts between levels of  $F$  are no bigger for  $\text{Nest}(\Delta, \Gamma)$  than for  $\text{Cross}(\Delta, \Gamma)$ .*

This is not quite the whole story, because it may be possible to construct a  $\text{Cross}(\Delta, \Gamma)$  for better block designs  $\Delta$  and  $\Gamma$  than a  $\text{Nest}(\Delta, \Gamma)$ .

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This all leads to interesting questions.  
Thank you for listening.