Design and analysis of experiments testing for biodiversity



Theoretical advances in experimental design Royal Statistical Society 2019 International Conference Belfast, September 2019

Ongoing joint work with Julia Reiss and Daniel Perkins

Bailey

Experiments in biodiversity

RSS, Belfast, 2019

It is now widely believed that biological diversity is good for the environment. One way that ecologists test this is to place random collections of species in mini-environments and then measure some outcome. Others use a carefully chosen collection of species.

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I have been working with a group of fresh-water ecologists on the design and analysis of such experiments. Our subsets of species are carefully chosen, not random. We design the combinations of these subsets with levels of other factors.

We also fit a nested family of plausible models to the data.

Our results suggest that the underlying model is not diversity at all. One of my crucial inputs has been the use of Hasse diagrams as a way of understanding a complicated family of plausible models for the expectation of the response.

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This seems to be the received wisdom.

Treatments:random sets of speciesMeasured response Y:some eco-desirable outcomeConclusion:the greater the number of different species,
the better the outcome.

A, B, C, D, E, F — six types of freshwater "shrimp". Put 12 shrimps in a jar containing stream water and alder leaf litter. Measure how much leaf litter is eaten after 28 days.

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identity			Level
A,, F	monoculture	12 of type A	1
AB,, EF	duoculture	6 of A, 6 of B	2
ABC, \ldots, DEF	triculture	4 of A, 4 of B, 4 of C	3

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20	ABC,, DEF	triculture	4 of A, 4 of B, 4 of C	3
11				

41

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The experiment was carried out in 4 blocks of 41 jars.

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$\frac{1}{41}$				

The experiment was carried out in 4 blocks of 41 jars. Actually 42 jars, because untreated jars were included, but their data was so obviously different that it was excluded from further modelling.

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Initial model fitting

The biologist fitted the model 'Richness' with 3 parameters, one for each level of richness, and found no evidence of any differences between the levels.

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and found no evidence of any differences between the levels.

This model for the response *Y* is

$$\mathbb{E}(Y) = \begin{cases} \alpha_1 & \text{on monocultures } A, \dots, F \\ \alpha_2 & \text{on duocultures } AB, \dots, EF \\ \alpha_3 & \text{on tricultures } ABC, \dots, DEF \end{cases}$$

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The data did not give any evidence against the null hypothesis that

 $\alpha_1 = \alpha_2 = \alpha_3$:

this is the 'Constant' model, or null model.

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Call in a statistician

	Assemblage identity			x1	<i>x</i> 2	<i>x</i> 3	<i>x</i> 4	<i>x</i> 5	<i>x</i> 6
1	Α	12 of type A	1	12	0	0	0	0	0
:			÷						
6	F	12 of type F	1	0	0	0	0	0	12
7	AB	6 of <i>A</i> , 6 of <i>B</i>	2	6	6	0	0	0	0
÷			:						
21	EF	6 of <i>E</i> , 6 of <i>F</i>	2	0	0	0	0	6	6
22	ABC	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>C</i>	3	4	4	4	0	0	0
÷			:						
41	DEF	4 of <i>D</i> , 4 of <i>E</i> , 4 of <i>F</i>	3	0	0	0	4	4	4

Call in a statistician

	Assemblage identity			<i>x</i> 1	<i>x</i> 2	х3	<i>x</i> 4	<i>x</i> 5	<i>x</i> 6
1	Α	12 of type A	1	12	0	0	0	0	0
:			÷						
6	F	12 of type F	1	0	0	0	0	0	12
7	AB	6 of <i>A</i> , 6 of <i>B</i>	2	6	6	0	0	0	0
:			÷						
21	EF	6 of <i>E</i> , 6 of <i>F</i>	2	0	0	0	0	6	6
22	ABC	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>C</i>	3	4	4	4	0	0	0
:			÷						
41	DEF	4 of <i>D</i> , 4 of <i>E</i> , 4 of <i>F</i>	3	0	0	0	4	4	4
I suggested the model 'Type' with 6 parameters β_1, \ldots, β_6 :									

$$\mathbb{E}(Y) = \sum_{i=1}^{6} \beta_i x_i$$

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:			:						
6	F	12 of type F	1	0	0	0	0	0	12
7	AB	6 of <i>A</i> , 6 of <i>B</i>	2	6	6	0	0	0	0
:			:						
21	EF	6 of <i>E</i> , 6 of <i>F</i>	2	0	0	0	0	6	6
22	ABC	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>C</i>	3	4	4	4	0	0	0
:			÷						
41	DEF	4 of <i>D</i> , 4 of <i>E</i> , 4 of <i>F</i>	3	0	0	0	4	4	4
I suggested the model 'Type' with 6 parameters β_1, \ldots, β_6 :									

$$\mathbb{E}(Y) = \sum_{i=1}^{6} \beta_i x_i \qquad (\sum x_i = 12 \text{ always, so no need for intercept.})$$

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Analysis of Variance (ANOVA) table

Source	df	SS	MS	F	Р
Richness	2	0.000009	0.000005	0.49	n.s.
Туре	5	0.003859	0.000772	81.37	< 0.0005
Richness * Type	10	0.000127	0.000013	1.34	n.s.
Assemblage Identity	23	0.000105	0.000005	0.48	n.s.
Block	3	0.000067	0.000022		
Error	120	0.001138	0.000009		
Total	163	0.005306			

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Verbatim from Journal of Animal Ecology

Scaling the Hasse diagram of expectation subspaces

```
Suppose that V_1 and V_2 are expectation subspaces,
with V_1 < V_2,
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Suppose that V_1 and V_2 are expectation subspaces, with V_1 < V_2, and an edge joining V_1 to V_2.
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The mean square for the extra fit in V_2 compared to the fit in V_1 is

 $\frac{\text{SS(fitted values in } V_2) - \text{SS(fitted values in } V_1)}{\dim(V_2) - \dim(V_1)}$

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Scale the Hasse diagram so that each edge has length proportional to the relevant mean square, and show the residual mean square to give a scale.

Assemblage ID Richness + Type Type Scale: $3 \times$ residual mean square **Richness** Constant

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Assemblage ID Richness + Type Type Conclusions: Scale: $3 \times$ residual mean square **Richness** Constant Experiments in biodiversity

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10/29

Assemblage ID Richness + Type Type Concluse Conclusions: The model Richness does not explain the data. The model Type explains the data well. Scale: $3 \times$ residual mean square **Richness** Experiments in biodiversity RSS, Belfast, 2019

10/29

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Assemblage ID Richness * Type Richness + Type Type Conclus Conclusions: The model Richness does not explain the data. The model Type explains the data well. There is no evidence that any larger model does any better. Scale: $3 \times$ residual mean square **Richness**

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Assemblage ID Richness + Type Type Concluse Conclusions: The model Richness does not explain the data. The model Type explains the data well. There is no evidence that any larger model does any better. Two experiments, with two responses each, all led to similar conclusions. Scale: $3 \times$ residual mean square Richness

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For complicated families of models, the biologists that I work with often find the Hasse diagram easier to understand than the equations.
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Similarly, they find the scaled Hasse diagram easier to understand than the anova table.

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63

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"Do I really need all 35 tricultures?"

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"Use 7 tricultures making a balanced incomplete-block design."

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Another success: *Advances in Ecological Research* published this picture of the Fano plane.

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How should such experiments be designed?

► To find out whether the response changes with different levels of richness?

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- To estimate the parameters (response per individual for each type) for the model Type?

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- To discriminate between the model Type and the more general model which allows the response of each type to depend on what other types are present?
- To estimate the parameters (response per individual for each type) for the model Type?
 - If so, we should not include any polycultures.

Suppose that there are t types in all. For a given level k of richness, each treatment consists of equal numbers n of each type in some subset of k types. Suppose that we can include b such treatments.

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Should we

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Should we

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- choose the subsets that would be best for an incomplete-block design with the usual linear model

(e.g. the Fano plane)? (*b* blocks, each of *k* plots; one type on each plot; each plot measured once)

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adopt some other strategy?





design	usual IBD model: 12 responses	
A 3/1	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	
$C \xrightarrow{2} B$		
$F \xrightarrow{6} 5 E$		
$F \xrightarrow{6} 1 B$		
5 2 F		
$\begin{array}{c} L \\ 4 \\ D \end{array}$	Experiments in biodiversity	RSS, Belfast, 2019

	design	usual IBD model: 12 responses	
	A 3/1	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	
С	$\frac{2}{D}B$	$ ext{Var}(\hat{ au}_A-\hat{ au}_B)=rac{4\sigma^2}{3}$	
F	6/5		
F			
5	2		
E	$4 \sqrt{3}$		
		Experiments in biodiversity	RSS, Belfast, 2019

_	design	usual IBD model: 12 responses	
	A 3/1	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	
С	$\frac{2}{D}B$	$ ext{Var}(\hat{ au}_A-\hat{ au}_B)=rac{4\sigma^2}{3}$	
F	6/4 5 E	$ au_A - au_D$ is not estimable	
F			
5 E			
	4 3 D	Experiments in biodiversity	RSS, Belfast, 2019

	design	usual IBD model: 12 responses	duocultures: 6 responses
	A 3/1	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	
С	$\frac{2}{D}B$	$ ext{Var}(\hat{ au}_A-\hat{ au}_B)=rac{4\sigma^2}{3}$	
F	6/5	$ au_A - au_D$ is not estimable	
F	A 6 1 B		
5	2		
E	4 3 C		
	D	Experiments in biodiversity	RSS, Belfast, 2019

design	usual IBD model: 12 responses	duocultures: 6 responses
A 3/1	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$
$C \frac{2}{D} B$	$\mathrm{Var}(\hat{ au}_A-\hat{ au}_B)=rac{4\sigma^2}{3}$	
$F \xrightarrow{6} 5 E$	$ au_A - au_D$ is not estimable	
F B		
5 2		
E 4 3 C		
D	Experiments in biodiversity	RSS. Belfast, 2019

	design	usual IBD model: 12 responses	duocultures: 6 responses
	A	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$
С	$\frac{2}{D}B$	$ ext{Var}(\hat{ au}_A - \hat{ au}_B) = rac{4\sigma^2}{3}$	$\hat{\beta}_A = \frac{Y_1 + Y_3 - Y_2}{2n}$
F	$\frac{6}{5}$	$ au_A - au_D$ is not estimable	
-	A		
F	β B		
5	2		
Ε	4 3 C		
	Ď	Experiments in biodiversity	RSS, Belfast, 2019

	design	usual IBD model: 12 responses	duocultures: 6 responses			
	A	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$			
С	$\frac{2}{D}B$	$ ext{Var}(\hat{ au}_A - \hat{ au}_B) = rac{4\sigma^2}{3}$	$\hat{eta}_A = rac{Y_1+Y_3-Y_2}{2n}$ $ ext{Var}(\hat{eta}_A) = rac{3\sigma^2}{4\sigma^2}$			
F	5 4	$ au_A - au_D$ is not estimable	$4n^2$			
F						
5 E	4 3 C					
	D	Experiments in biodiversity	RSS, Belfast, 2019			

	design	usual IBD model: 12 responses	duocultures: 6 responses			
	A	$\mathbb{E}(\Upsilon_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$			
С	$\frac{3}{2}$ B	$\operatorname{Var}(\hat{ au}_A - \hat{ au}_B) = rac{4\sigma^2}{3}$	$\hat{eta}_A = rac{Y_1+Y_3-Y_2}{2n} onumber Var(\hat{eta}_A) = rac{3\sigma^2}{4\pi^2}$			
F	5 4 E	$ au_A - au_D$ is not estimable	$\sum_{i} \operatorname{Var}(\hat{\beta}_i) \text{ is minimized}$			
F						
э Е						
		Experiments in biodiversity	RSS, Belfast, 2019			

	design	usual IBD model: 12 responses	duocultures: 6 responses
	A	$\mathbb{E}(\Upsilon_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$
C F	3 2 B D B 6 4 5 E	$ ext{Var}(\hat{ au}_A - \hat{ au}_B) = rac{4\sigma^2}{3}$ $ au_A - au_D$ is not estimable	$\hat{\beta}_{A} = \frac{Y_{1} + Y_{3} - Y_{2}}{2n}$ $\operatorname{Var}(\hat{\beta}_{A}) = \frac{3\sigma^{2}}{4n^{2}}$ $\sum \operatorname{Var}(\hat{\beta}_{i}) \text{ is minimized}$
	<u></u>	2	i
	$6 \bigwedge^{\Lambda} 1$	$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_D) = 3\sigma^2$	
F 5		$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_C) = \frac{8\sigma^2}{3}$	
E	4 3	$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_B) = \frac{5b^2}{3}$	
	D	Experiments in biodiversity	RSS, Belfast, 2019

_	design	usual IBD model: 12 responses	duocultures: 6 responses			
	A	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$			
С	$\frac{5}{2} B$	$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_B) = \frac{4\sigma^2}{3}$	$\hat{eta}_A = rac{Y_1+Y_3-Y_2}{2n} onumber \ ext{Var}(\hat{eta}_A) = rac{3\sigma^2}{4n^2}$			
F	5 E	$ au_A - au_D$ is not estimable	$\sum_{i} \operatorname{Var}(\hat{\beta}_i) \text{ is minimized}$			
-	A	$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_D) = 3\sigma^2$				
F 5		$\operatorname{Var}(\hat{ au}_A - \hat{ au}_C) = rac{8\sigma^2}{3}$				
E		$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_B) = \frac{5\sigma^2}{3}$				
	$4 \lor 3$ D	$\sum \operatorname{Var}(\hat{\tau}_i - \hat{\tau}_j)$ is minimized				
	-	$i \neq j$ Experiments in biodiversity	RSS, Belfast, 2019			

	design	usual IBD model: 12 responses	duocultures: 6 responses
	A	$\mathbb{E}(Y_{1A}) = \tau_A + \delta_1$	$\mathbb{E}(Y_1) = n(\beta_A + \beta_B)$
C F	$ \begin{array}{c} 3 \\ 2 \\ D \\ 6 \\ 4 \\ 5 \\ E \end{array} $	$\operatorname{Var}(\hat{ au}_A - \hat{ au}_B) = rac{4\sigma^2}{3}$ $ au_A - au_D$ is not estimable	$\hat{\beta}_A = \frac{Y_1 + Y_3 - Y_2}{2n}$ $\operatorname{Var}(\hat{\beta}_A) = \frac{3\sigma^2}{4n^2}$ $\sum_i \operatorname{Var}(\hat{\beta}_i) \text{ is minimized}$
-	A	$\operatorname{Var}(\hat{\tau}_A - \hat{\tau}_D) = 3\sigma^2$	1
F 5 E		$\operatorname{Var}(\hat{ au}_A - \hat{ au}_C) = rac{8\sigma^2}{3}$ $\operatorname{Var}(\hat{ au}_A - \hat{ au}_B) = rac{5\sigma^2}{3}$	β_A is not estimable
	4 3 D	$\sum_{i \neq j} \operatorname{Var}(\hat{\tau}_i - \hat{\tau}_j) \text{ is minimized}$ Experiments in biodiversity	RSS, Belfast, 2019

Consider incomplete-block designs for *t* treatments in *b* blocks of size *k*.

Usual model expected response on any unit with treatment *i* in block *B* is $\tau_i + \delta_B$

Polyculture model expected response on any unit with a collection of *n* individuals from each species in subset *B* is $n \sum_{i \in B} \beta_i$

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If there is no balanced incomplete-block design for t treatments in b blocks of size k then a design which is best for one situation may be worst for the other.

One aspect of a third biodiversity experiment: 4 types of "shrimp"

		(Composition	Richness	<i>x</i> 1	<i>x</i> 2	<i>x</i> 3	<i>x</i> 4
	1	Α	12 of type A	1	12	0	0	0
	2	В	12 of type B	1	0	12	0	0
	3	С	12 of type C	1	0	0	12	0
	4	D	12 of type D	1	0	0	0	12
	5	AB	6 of <i>A</i> , 6 of <i>B</i>	2	6	6	0	0
	6	AC	6 of <i>A</i> , 6 of <i>C</i>	2	6	0	6	0
	7	AD	6 of <i>A</i> , 6 of <i>D</i>	2	6	0	0	6
	8	BC	6 of <i>B</i> , 6 of <i>C</i>	2	0	6	6	0
	9	BD	6 of <i>B</i> , 6 of <i>D</i>	2	0	6	0	6
	10	CD	6 of <i>C</i> , 6 of <i>D</i>	2	0	0	6	6
	11	ABC	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>C</i>	3	4	4	4	0
	12	ABD	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>D</i>	3	4	4	0	4
	13	ACD	4 of <i>A</i> , 4 of <i>C</i> , 4 of <i>D</i>	3	4	0	4	4
	14	BCD	4 of <i>B</i> , 4 of <i>C</i> , 4 of <i>D</i>	3	0	4	4	4
	15	ABCD	3 each of <i>A</i> , <i>B</i> , <i>C</i> and <i>D</i>	4	3	3	3	3
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Family of expectation models (so far)


Family of expectation models (so far)



Five responses were measured. For every response, the sum of squares of fitted values for Composition was hardly any bigger than the sum of squares of fitted values for the model Richness * Type, so we decided to omit Richness * Type.

Each of the 15 compositions was combined with three temperatures: 5° C, 10° C and 15° C.

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Each of the 45 combinations was replicated twice.

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Each of the 45 combinations was replicated twice.

Three temperature-controlled rooms in a lab were used. Each room had a single temperature and two of each composition. Therefore there was no appropriate residual mean square to compare the main effect of Temperature with, but all other effects could be assessed.

Diagram from a paper in Global Change Biology



Bailey

20/29

For each single type of response, Type \times Temperature explained the data well, with no need for further terms.

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Note that this is a simple consequence of the model

 $\beta_1 x_1 + \beta_2 x_2 + \beta_3 x_3 + \beta_4 x_4$

if the rankings of β_1 , β_2 , β_3 and β_4 are different over the five types of response.

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One aspect of a fourth biodiversity experiment

A, *B*, *C*— types of freshwater "shrimp". Put 12 shrimps in a jar with stream water and alder leaf litter. Measure how much leaf litter is eaten after 28 days.

Experimental unit = jar.

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Assemblage identity			Richness	x1	<i>x</i> 2	х3
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4	AB	6 of <i>A</i> , 6 of <i>B</i>	2	6	6	0
5	AC	6 of <i>A</i> , 6 of <i>C</i>	2	6	0	6
6	BC	6 of <i>B</i> , 6 of <i>C</i>	2	0	6	6
7	ABC	4 of <i>A</i> , 4 of <i>B</i> , 4 of <i>C</i>	3	4	4	4

Family of expectation models (so far)



Family of expectation models (so far)



For these numbers, Assemblage identity = Richness * Type.

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Experiments in biodiversity

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The other aspect of the biodiversity experiment



Fig 1. Photographs of the structures used to create habitat complexity in microcosms with 'structure present'. The basic unit of each structure was a plastic plant strip (mimicking *Ceratophyllum* sp.), joined up as a ring (- 8cm in diameter) and four levels of fractal dimension were created with them: 1) level 1 consisted of two rings aligned, with a fractal dimension (D) of 1.77; 2) level 2 consisted of two rings twisted into each other (D = 1.80); 3) level 3 consisted of three rings locked together (D = 1.81) and 4) level four was a ball made from 3 rings together (D = 1.80). This design threefore also gave two levels of 'amount of structure' - 3 Experiments' in block urcerster and 4.5 g for complexity level 3 and 4. RSS, Belfast, 2019

Hasse diagram for environmental model subspaces



The experiment: 3 blocks, each with 35 jars

Environment	Assemblage identity						
Complexity	A	В	C	AB	AC	BC	ABC
0	×	×	×	×	×	×	\times
1	×	×	×	×	×	×	\times
2	×	×	\times	\times	×	×	\times
3	×	×	\times	\times	×	×	\times
4	×	×	×	\times	×	×	\times

The experiment: 3 blocks, each with 35 jars

Environment	Assemblage identity						
Complexity	A	В	С	AB	AC	BC	ABC
0	×	×	×	×	×	×	\times
1	×	×	×	×	×	×	×
2	\times	\times	\times	\times	×	×	\times
3	\times	×	\times	\times	×	×	×
4	×	×	×	×	×	×	×

Spanish PhD student Lorea Flores visited the University of Roehampton for three months;

gathered the "shrimps" from ponds on the campus;

put the combinations of leaves, shrimps and plastic rings into jars;

put one jar of each type onto each of three shelves in a temperature-controlled room;

measured various responses on each jar (some daily, some at the end).

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Experiments in biodiversity

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The models consist of all interactions and sums of those shown in the two previous diagrams (the gentle reader can draw her own Hasse diagram!).

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... but the ecologists cannot do this. They can use statistical software to fit each model, and then use a spreadsheet to subtract sums of squares appropriately. This is error-prone.

Solution! Summer student Justin Thong dug into the statistical software R to find a short sequence of commands that gives precisely the right output (not straightforward, because R makes some stupid assumptions).

Bailey

Experiments in biodiversity

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RESEARCH ARTICLE

Habitat Complexity in Aquatic Microcosms Affects Processes Driven by Detritivores

Lorea Flores¹*, R. A. Bailey^{2,3}, Arturo Elosegi⁴, Aitor Larrañaga⁴, Julia Reiss⁵*

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Abstract

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Individual species numbers; Plastic rings or not; Number of plastic rings.

Nothing more complicated, so not Richness, not Fractal dimension, no interactions.