



## Introduction of the design key

H. D. Patterson: The factorial combination of treatments in rotation experiments. *Journal of Agricultural Science*, **65** (1965), 171–182.

This paper introduced the design key.

The number of levels of each factor must be a power of a single prime number p. All examples have p = 2, but it is mentioned that the method can also be used with p = 3.

I had intended to include an example from this paper, but they are all far too complicated. I am amazed that any referee understood it at the time. Parts of it are almost as if Desmond is thinking aloud.

| ample 1 (Graeco-       | Latin square): the        | key                      | Example 1 (Graeco-Latin square): construction  |  |  |  |  |
|------------------------|---------------------------|--------------------------|--|--|--|--|--|
| Factors                | 7                         |                          |  |  |  |  |  |
| <b>T</b> 1 1 1         | Factors                   | Factors with five levels | $W = R + C \qquad N = R + 2C$  |  |  |  |  |
| Experimental units     | Rows                      | ĸ                        |  |  |  |  |  |
|                        | Columns                   | С                        | С  |  |  |  |  |
| Treatments             | Variety of wheat          | W                        | $R \mid 0 \mid 1 \mid 2 \mid 3 \mid 4 \mid$  |  |  |  |  |
|                        | Quantity of nitrogen      | Ν                        | 0 0,0 1,2 2,4 3,1 4,3  |  |  |  |  |
| Every factor is repre  | sented by a single lette  | r                        |  |  |  |  |  |
| Levels are integers n  | odulo 5                   |                          |  |  |  |  |  |
| Levels are integers in | iodulo 5.                 |                          | $\frac{2}{3}$ $\frac{2}{3}$ $\frac{2}{3}$ $\frac{2}{4}$ $\frac{1}{0}$ $\frac{1}{0}$ $\frac{1}{2}$ $\frac{1}{4}$ $\frac{1}{2}$ $\frac{1}{1}$  |  |  |  |  |
| Constraints The trea   | tment factors W and N     | should both be           | $\frac{3}{4} \frac{3}{4} \frac{3}$ |  |  |  |  |
| orthogonal to rows a   | ind to columns.           |                          | 4 4,4 0,1 1,5 2,0 5,2  |  |  |  |  |
| Design key The desi    | ion key eynresses each    | treatment factor as a    | The experimental units are defined by all combinations of  |  |  |  |  |
| linear combination of  | factors on the experiment | nontal units             | levels of R and C  |  |  |  |  |
|                        | i lactors on the experin  | nental units.            | levels of K and C.   |  |  |  |  |
| W                      | = R + C $N = R + C$       | 2 <i>C</i>               | The level of <i>W</i> is shown first in each cell.   |  |  |  |  |
|                        |                           |                          | The level of N is shown second in each cell.   |  |  |  |  |
|                        |                           |                          |  |  |  |  |  |
|                        |                           | 5/33                     | N  |  |  |  |  |

| Example 1 (Graeco-Latin square): confounding  |  |                                       |  |   | Treatment factors  |  |  |  |
|---|--|---------------------------------------|--|---|--|--|--|--|
| Ν   | V = R + C  | Ν                                     | = R + 2C   |   | There is a set $\mathcal{F}$ of treatment factors.<br>There is one potential treatment for each combination of levels of all the factors in $\mathcal{F}$ .  |  |  |  |
| Stratum<br>Rows<br>Columns<br>Rows-by-Columns | unit effect<br>R<br>C<br>R + C<br>R + 2C<br>R + 3C<br>R + 4C | df<br>4<br>4<br>4<br>4<br>4<br>4<br>4 | $\frac{\text{tmt factor}}{W + 2N}$ $\frac{W + 4N}{W}$ $\frac{W}{N}$ $W + 3N$ $W + N$ | tmt effect<br>interaction<br>interaction<br>variety main<br>nitrogen main<br>interaction<br>interaction | At first, we assume that every factor in $\mathcal{F}$ has $p$ levels,<br>where $p$ is prime. The levels are the integers modulo $p$ .<br>All addition is done modulo $p$ .<br>Each non-zero linear combination of factors in $\mathcal{F}$ gives a<br>treatment pseudofactor with $p$ levels.<br>This gives $p - 1$ degrees of freedom for contrasts between<br>treatments, all belonging to the interaction of those genuine<br>treatment factors whose coefficient is non-zero. |  |  |  |
| W + N   | $= 2R + 3C \equiv$   | ≡ 6R                                  | +9C = R +  | - 4C  | If one such linear combination is a non-zero multiple of<br>another, then they correspond to the same df; otherwise the<br>corresponding sets of contrasts are orthogonal to each other.   |  |  |  |

| Factors on the experimental units   | Poset block structure in Example 1   |
|---|--|
| <ul> <li>There is a set G of unit factors.</li> <li>We assume that the real factors on the experimental units form a poset block structure.</li> <li>This means that they can be defined by a panel diagram, showing <ul> <li>the list of factors G<sub>1</sub>,, G<sub>m</sub> in G</li> <li>for each G<sub>i</sub>, its number n<sub>i</sub> of levels;</li> <li>for each G<sub>i</sub>, what it is nested in.</li> </ul> </li> <li>There are n<sub>1</sub> × ··· × n<sub>m</sub> experimental units, one for each combination of levels of G<sub>1</sub>,, G<sub>m</sub>.</li> <li>"G<sub>i</sub> is nested in G<sub>j</sub>" means</li> </ul> | <ul> <li>5 Rows</li> <li>5 Columns</li> </ul> This panel diagram tells us that <ul> <li>there are factors <i>R</i> and <i>C</i>, each with 5 levels;</li> <li>there are 25 experimental units, one for each combination of levels of <i>R</i> and <i>C</i>;</li> <li>there is no nesting;</li> <li>the real factors on the experimental units are</li> </ul> |
| "if two objects have the same level of $G_i$ then this has<br>no significance unless they have the same level of $G_j$ ".<br>The real factors are combinations of levels of none or more of   |  |
| $G_1, \ldots, G_m$ subject to the rule that<br>if $G_i$ is included and $G_i$ is nested in $G_j$ then $G_j$ must be included.   | <i>RC</i> with 25 levels.  |

| Poset block structure in Example 2  | Powers of a prime  |  |  |  |  |
|---|--|--|--|--|--|
| 4 Blocks<br>4 Plots in B  |  |  |  |  |  |
| This panel diagram tells us that  | If a factor has $p^r$ levels, where $r \ge 2$ ,  |  |  |  |  |
| there are factors B and P, each with 4 levels;  | then it is represented by <i>r</i> pseudofactors, each with <i>p</i> levels.                                       |  |  |  |  |
| <ul> <li>there are 16 experimental units,<br/>one for each combination of levels of <i>B</i> and <i>P</i>;</li> </ul>   | The convention is that these pseudofactors are written with the same single letter and subscripts $1, \ldots, r$ . |  |  |  |  |
| <ul> <li><i>P</i> is nested in <i>B</i>,<br/>so there is no real factor involving <i>P</i> but not <i>B</i>;</li> </ul> |  |  |  |  |  |
| the real factors on the experimental units are  |  |  |  |  |  |
| $\oslash$ with 1 level;   |  |  |  |  |  |
| B with 4 levels;  |  |  |  |  |  |
| <i>BP</i> with 16 levels.   |  |  |  |  |  |
| 11/3  | 12/3   |  |  |  |  |

| Identification of factorial effects  | Example 2 (Factorial design in blocks): the key  |
|--|--|
| <ul> <li>For a linear combination of factors (and pseudofactors) in <i>F</i> or a linear combination of factors (and pseudofactors) in <i>G</i>, we need to identify the factorial effect containing the corresponding <i>p</i> – 1 degrees of freedom.</li> <li>1. Write down all the letters which occur, ignoring subscripts;</li> <li>2. if factor <i>C</i> is nested in factor <i>D</i> and letter <i>C</i> occurs then include letter <i>D</i>;</li> <li>3. remove any duplicate letters.</li> <li>The set of letters remaining gives the factorial effect.</li> </ul> | FactorsFactorsWith 2 levelsExperimental unitsBlocks (4) $B_1, B_2$ Plots in Blocks (4) $P_1, P_2$ Treatments $S$ (2) $S$ T (2)TU (2)UV (2)VThe pseudofactors for each factor all have the same letter.Levels are integers modulo 2.Constraints All treatment main effects should be orthogonal to blocks. So should as many two-factor interactions as possible.Design key $S = P_1$ $T = P_2$ $U = B_1 + P_1 + P_2$ $V = B_2 + P_1 + P_2$ |

| Example 2 (Factorial design in blocks): construction  | Example 2 (Factorial design in blocks): confounding             |  |  |  |  |
|---|---|--|--|--|--|
|   | $S = P_1$ $T = P_2$ $U = B_1 + P_1 + P_2$ $V = B_2 + P_1 + P_2$ |  |  |  |  |
| $S = P_1$ $T = P_2$ $II = R_1 + P_1 + P_2$ $V = R_2 + P_1 + P_2$  | Stratum unit effect df tmt factor tmt effect                    |  |  |  |  |
| $b = r_1$ $r_1 = r_2$ $a = b_1 + r_1 + r_2$ $v = b_2 + r_1 + r_2$   | Blocks (B) $B_1$ 1 $S+T+U$ 3 f.i.                               |  |  |  |  |
| The experimental units are defined by all combinations of   | $B_2$ 1 $S + T + V$ 3 f.i.                                      |  |  |  |  |
| levels of $B_1$ , $B_2$ , $P_3$ and $P_2$   | $B_1 + B_2$ 1 $U + V$ $U$ -by- $V$ intr                         |  |  |  |  |
|   | Plots in Blocks (BP) $P_1$ 1 S main S                           |  |  |  |  |
|   | $P_2$ 1 T main T  |  |  |  |  |
| Block 1 Block 2 Block 3 Block 4   | $P_1 + P_2$ 1 $S + T$ S-by-T intr                               |  |  |  |  |
| $ \begin{array}{c c c c c c c c c c c c c c c c c c c $   | $B_1 + P_1$ 1 $T + U$ $T$ -by- $U$ intr                         |  |  |  |  |
| $B_2 \   \ 0 \   \ 0 \   \ 0 \   \ 0 \   \ 1 \   \ 1 \   \ 1 \   \ 1 \   \ 0 \   \ 0 \   \ 0 \   \ 0 \   \ 1 \ 1$ | $B_1 + P_2$ 1 $S + U$ $S$ -by- $U$ intr                         |  |  |  |  |
| $P_1 \mid 0 \mid 0 \mid 1 \mid 1 \mid 0 \mid 0 \mid 1 \mid 1 \mid 0 \mid 0$   | $B_1 + P_1 + P_2  1 \qquad U \qquad \text{main } U$             |  |  |  |  |
| $P_2 \mid 0 \mid 1 \mid 0 \mid 1$   | $B_2 + P_1$ 1 $T + V$ $T$ -by- $V$ intr                         |  |  |  |  |
| S         0         0         1         1         0         0         1         1         0         0         1         1         0         0         1         1   | $B_2 + P_2$ 1 $S + V$ S-by-V intr                               |  |  |  |  |
| $T \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \$   | $B_2 + P_1 + P_2  1 \qquad V \qquad \text{main } V$             |  |  |  |  |
| $egin{array}{c c c c c c c c c c c c c c c c c c c $  | $B_1 + B_2 + P_1$ 1 $S + U + V$ 3 f.i.                          |  |  |  |  |
| $V \parallel 0 \mid 1 \mid 1 \mid 0 \mid 1 \mid 0 \mid 0 \mid 1 \mid 0 \mid 0$  | $B_1 + B_2 + P_2$ 1 $T + U + V$ 3 f.i.                          |  |  |  |  |
|   | $B_1 + B_2 + P_1 + P_2$ 1 $S + T + U + V$ 4 f.i.                |  |  |  |  |
| 15/33   | 16/3  |  |  |  |  |

| What does the design key do?  | Generalization to more than one prime  |
|---|--|
| <ul> <li>A design key is a list<br/>giving an alias for each treatment (pseudo-)factor as<br/>a linear combination of (pseudo-)factors for the experimental<br/>units.</li> <li>This gives <ul> <li>an algorithm for constructing the design;</li> <li>a design that is orthogonal;</li> <li>(if it is a fractional replicate) a fraction which is regular;</li> <li>an algorithm for identifying the confounding between<br/>treatment effects and strata defined by a poset block<br/>structure on the experimental units.</li> </ul> </li> </ul> | Suppose that more than one prime is involved.<br>If any factor has a composite number of levels, express it as a<br>product of pseudofactors, each with a prime number of levels.<br>For each prime $p_i$ separately, consider only the treatment<br>(pseudo-)factors and unit (pseudo-)factors which have $p_i$ levels,<br>and make a design key for them, using arithmetic modulo $p_i$ .<br>Suppose that $p_1, \ldots, p_k$ are among the primes involved, and<br>that, for $i = 1, \ldots, k$ , $T_i$ is a linear combination of treatment<br>factors or pseudofactors with $p_i$ levels.<br>Then $T_i$ belongs to an effect defined by a subset $S_i$ of the initial<br>letters of the genuine treatment factors.<br>It can be shown that the $\prod_{i=1}^{k} (p_i - 1)$ df for the interaction<br>between $T_1, \ldots, T_k$ all belong to the effect defined by the subset<br>$S_1 \cup \cdots \cup S_k$ .<br>The analogous result holds for factors on the experimental<br>units. |



| Example 3 | Example <u>3</u> (Whole-plot factors): skeleton anova  |  |  |   |       | Two-phase experiments   |  |  |  |  |
|-----------|--|--|--|---|-------|---|--|--|--|--|
|           | units<br>source<br>Mean<br>Rows<br>Columns<br>R#C<br>Strips[R]<br>Lines[C]<br>Strips[R]#C<br>R#Lines[C]<br>S[R]#L[C] | df           1           3           3           4           8           12           8           32 | treatments<br>source<br>Mean<br>Grass type<br>Residual<br>Mowing height<br>Residual<br>Fertilizer quantity<br>Residual<br>G#M<br>Residual<br>G#F<br>Residual<br>M#F<br>G#M#F<br>Residual | df           1           2           2           2           2           2           2           2           2           2           6           2           6           4           4           24 | va. i | Treatments Phase I units Phase II units<br>The first design key allocates treatments to Phase I units.<br>The second design key allocates Phase I units to Phase II units.<br>Combining these allows us to keep track of confounding all the<br>way through,<br>which helps us to choose suitable design keys in the first place. |  |  |  |  |
|           |  |  |  |   | 21/33 | 24/3  |  |  |  |  |

| Example 4 (Proteomics): constraints   | Example 4 (Proteomics): confounding  |
|---|--|
| 2 Interventions<br>2 Interventions<br>2 Animals in C<br>2 Tissues + 2 Positions in A, C<br>2 Animals in C, + 2 Positions in A, C  | 2 Interventions<br>2 Tissues + 8 Cages<br>2 Animals in C<br>+ 2 Positions in A, C<br>8 Runs<br>8 Labels  |
| <ul> <li>Interventions probably has the biggest variance from<br/>Phase I, so try to confound this with a low-variance term in<br/>Phase II.</li> <li>If possible, confound the rest of Cages with the same term,<br/>to avoid losing degrees of freedom for the residual.</li> <li>If possible, make Tissues and I#T orthogonal to Runs and<br/>Labels.</li> </ul> | $I = C_1 	 C_1, 	 C_2, 	 C_3 	 C_i = R_i + L_i 	 R_1, 	 R_2, 	 R_3 	 A 	 A = R_1 	 T 	 T = P + C_3 	 P 	 P = L_2 	 L_1, 	 L_2, 	 L_3 	 P + C_2 = R_2 	 Positions[A,C], 	 Runs$ |
| <b>Design key</b><br>$L = L = C_1 + C_2 + C_2 + C_2 + L_1 + R_2 + R_2 + R_3$  | $T = P + C_3 = L_2 + R_3 + L_3$ T, P[A,C], R#L   |
| $\begin{bmatrix} T & T = C_1 & C_1, & C_2, & C_3 & C_1 - R_1 + L_1 & R_1, & R_2, & R_3 \\ A & A = R_1 \\ T & T = P + C_3 & P & P = L_2 & L_1, & L_2, & L_3 \end{bmatrix}$   | $I + T = C_1 + P + C_3 = R_1 + L_1 + L_2 + R_3 + L_3$ I#T, P[A,C], R#L   |

| units  |    | animal-bits                 | treatments |               |    |  |
|--------|----|-----------------------------|------------|---------------|----|--|
| source | df | source                      | df         | source        | df | EMS  |
| Mean   | 1  | Mean                        | 1          | Mean          | 1  | $\xi_0 + 2\eta_0 + q_0$                      |
| Runs   | 7  | Animals[C] <sub>1</sub>     | 1          |               |    | $\xi_{\rm R} + 2\eta_{\rm CA}$               |
|        |    | Positions[A,C] <sub>1</sub> | 2          |               |    | $\xi_{\rm R} + 2\eta_{\rm CAP}$              |
|        |    | Residual                    | 4          |               |    | ξ̃r  |
| Labels | 7  | Animals[C] <sub>2</sub>     | 1          |               |    | $\xi_{\rm L} + 2\eta_{\rm CA}$               |
|        |    | Positions[A,C] <sub>2</sub> | 2          |               |    | $\xi_{\rm L} + 2\eta_{\rm CAP}$              |
|        |    | Residual                    | 4          |               |    | $\xi_{\rm L}$                                |
| R#L    | 49 | Cages                       | 7          | Interventions | 1  | $\xi_{\rm RL} + 2\eta_{\rm C} + q({\rm I})$  |
|        |    |                             |            | Residual      | 6  | $\xi_{\rm RL} + 2\eta_{\rm C}$               |
|        |    | Animals[C] <sub>3</sub>     | 6          |               |    | $\xi_{\rm RL} + 2\eta_{\rm CA}$              |
|        |    | Positions[A,C] <sub>3</sub> | 12         | Tissues       | 1  | $\xi_{\rm RL} + 2\eta_{\rm CAP} + q({\rm T}$ |
|        |    |                             |            | I#T           | 1  | $\xi_{\rm RL} + 2\eta_{\rm CAP} + q(\Gamma)$ |
|        |    |                             |            | Residual      | 10 | $\xi_{\rm RL} + 2\eta_{\rm CAP}$             |
|        |    | Residual                    | 24         |               |    | ξ̃rl   |

## Example 5 (Field then laboratory): constraints



$$V_1 = P_1$$
  $C$   
 $V_2 = P_2$   $P_1$ ,  $P_2$   $S_1$ ,  $S_2$ 

**Constraints** All Variety effects should be orthogonal to Rows and orthogonal to Columns in Phase I.

Then at least 2df for Varieties must be confounded with R#C, so there is no loss of generality in taking this design key for the first phase.

**Question** What should we do in the second phase, given that at least 2df for Varieties must be confounded with Batches?

| Example 5 (Fie   | ld then labora                        | atory): option   | 1   | Example 5 (Field then laboratory): option 2   |              |  |  |  |  |
|--|---------------------------------------|--|---|---|--------------|--|--|--|--|
| 27 Varieties<br>V <sub>1</sub> , V <sub>2</sub> , V <sub>3</sub> | $V_3 = R + C$ $V_1 = P_1$ $V_2 = P_2$ | $\begin{array}{c} \text{vs} \\ \text{umms} \\ \text{ts in R, C} \end{array}$ $\begin{array}{c} R \\ C \\ P_1, P_2 \end{array} P_i$ | 9 Batches<br>9 Samples in B<br>= $B_1$ $B_1$ , $B_2$<br>= $B_2$<br>= $S_i$ $S_1$ , $S_2$                          | 27 Varieties3 Rows<br>3 Columns<br>9 Plots in R, C9 Batches $V_1, V_2, V_3$ $V_3 = R + C$ $R$ $R = B_1$ $B_1, B_2$ $V_1 = P_1$ $C$ $P_1 = B_2$ $V_2 = P_2$ $P_1, P_2$ $C = S_1, P_2 = S_2$ $S_1, S_2$ |              |  |  |  |  |
| samples  | plots                                 | varieties  |   | samples   | plots        | varieties  |  |  |  |
| source df  | source df                             | source df  | EMS   | source d  | lf source di | source df  | EMS  |  |  |
| Batches 8  | Rows2Columns2R#C4                     | $V_3$ 2<br>Posidual 2  | $ \begin{array}{c} \zeta_{B} + \eta_{R} \\ \zeta_{B} + \eta_{C} \\ \zeta_{B} + \eta_{RC} + q(V_{3}) \end{array} $ | Batches   |              | $\begin{array}{c c} V_1 & 2\\ \hline Residual & 4 \end{array}$ | $\begin{aligned} \tilde{\xi}_{\rm B} &+ \eta_{\rm R} \\ \boldsymbol{\xi}_{\rm B} &+ \eta_{\rm RCP} + q(V_1) \\ \boldsymbol{\xi}_{\rm B} &+ \eta_{\rm RCP} \end{aligned}$ |  |  |
|  |                                       | Kesiuuai 2   | $\varsigma_{\rm B} \pm \eta_{\rm RC}$   |   |              |  |  |  |  |





